

# Government Revenue, Expenditure, and Public Debt in Mali: Empirical Evidence and Policy Implications

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## Abstract

This study examines the relationship between government revenue, expenditure, and public debt in Mali (2000–2024) using a Vector Error Correction Model (VECM). The Johansen cointegration test confirms a long-term equilibrium, highlighting structural fiscal imbalances. Results show that expenditure negatively impacts revenue (-1.3017 elasticity), indicating that excessive spending weakens revenue mobilization. Public debt has no significant effect on revenue, suggesting inefficiencies in debt-financed policies. In the short run, expenditure adjusts significantly to disequilibrium (ECT = 0.6998,  $p = 0.001$ ), while revenue and debt remain unresponsive. Furthermore, Granger causality tests, based on the Toda-Yamamoto approach confirms a bidirectional relationship between government revenue and expenditure, supporting both the tax-spend and spend-tax hypotheses. Additionally, there is supporting evidence that public debt is Granger-caused by both revenue and expenditure, indicating that fiscal deficits in Mali are primarily financed through borrowing. These findings stress the need for fiscal reforms in Mali, focusing on tax efficiency, prudent expenditure, and sustainable debt management to prevent macroeconomic instability.

**Keywords:** Fiscal sustainability, Revenue-expenditure nexus, Public debt dynamics, Mali

## Introduction

Characterized by agriculture, extractive industries, and reliance on external aid, Mali's economic framework is heavily contingent upon gold mining, rendering it vulnerable to global price volatilities (Mainguy, 2011; Drakenberg, 2010). The expansion of artisanal and small-scale mining (ASM) further complicates fiscal governance, as the sector's predominantly informal nature hinders effective revenue

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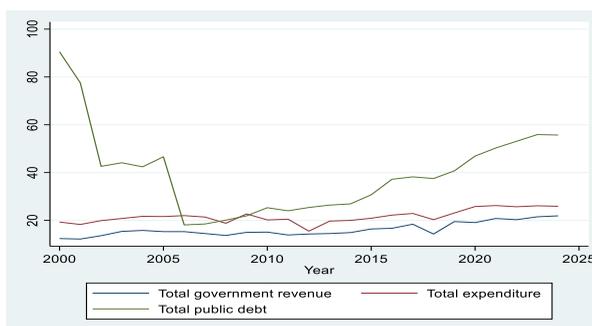
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collection (Traoré et al.; World Bank, 2023). Consequently, revenue mobilization is constrained by a fiscal policy encumbered by a limited tax base and a pervasive informal economy (Haavik & Cissé, 2024; Bah, 2024). Institutional and structural impediments inhibit sector formalization (Simba et al., 2023), despite promising strides in tax reform and digital revenue collection, which necessitate enhanced institutional backing (Mansour & Keen, 2009; Cogneau & Mo, 2024).

Mali grapples with a complex fiscal dichotomy: pursuing immediate developmental imperatives while ensuring long-term sustainability. Investment in social infrastructure is critically necessary yet curtailed by budget earmarking, endemic insecurity, and escalating military expenditures (Herrera & Ouedraogo, 2018; Guindo & Hak, 2024, ICG, 2024). In this precarious context, the increasing reliance on debt-financed public spending for development and deficit management heightens the risk of fiscal destabilization if not judiciously managed (Atingi-Ego et al., 2021; Shah et al., 2024). While strategically deployed debt can catalyze growth, rising global interest rates and commodity price fluctuations threaten sustainable fiscal management (AfDB, 2023, Nurjihadi & Guindo, 2024). Effective debt management, transparent governance, and access to concessional financing are imperative for maintaining fiscal stability (World Bank, 2023). Moreover, the adverse impacts of climate change exacerbate fiscal vulnerabilities, undermining agricultural productivity and exacerbating food insecurity (FAO, 2012). The fiscal burden of climate adaptation intensifies pressure on national resources, underscoring the need for substantial international support (AfDB, 2023).

**Figure 1.**

*The composition of public debt, government expenditure and revenue in Mali 2000-2024 (as % of GDP)*



Source: Author

This (**Error! Reference source not found.**) represents the trends in Total Government Revenue, Total Expenditures, and Total Public Debt in Mali over a 25-year period from 2000 to 2025. These trends tell us much about the fiscal and economic landscape in Mali, both progressing and with problems. Government revenue growth was modest, spurred on by agriculture, gold mining, and tax collection. Additionally, it is very vulnerable to external shocks (Staatz et al. 2011; Drakenberg 2010). After 2010, tax reforms coupled with mining expansion increased revenue collection, but the tax base continued to be narrow (Mainguy 2011; Thomas 2010). Meanwhile, expenditures grew at an increasing rate, with investments in infrastructure, social services, and security operations, often overtaking revenues, leaving deficits (Blimpo & Harding 2013; Maïga et al. 2021). In terms of public debt, it notably rose after 2010, being predominantly externally financed, bringing into question sustainability due to Mali's fragile revenue capacity (Atingi-Ego et al. 2021; Maddah et al. 2024). This debt accumulation endangers the priority public investment programs and increases fiscal vulnerability.

## Review of the Literature

### ❖ Theoretical Literature

Public debt, government expenditure, and revenue are at the center of public finance debates based upon differing theoretical perspectives. The tax smoothing theory (Barro 1979) argues that taxpayers sustain stable rates by borrowing during periods of increasing government expenditure to smooth out economic distortions. The displacement effect (Peacock & Wiseman 1961) contradicts this view, arguing that government expenditure increases in times of crises, leading to higher taxes and a sustained public debt rate. This puts the former hypothesis into the framework of the very same spending-led fiscal expansion that supports the spend-tax hypothesis.

The spend-tax hypothesis postulates, with the support of Barro's 1974 Ricardian equivalence, a direct relationship of causation between government spending, revenues, and debts, asserting that a government may either subject itself to taxation or borrow money: the two acts are interchangeable. The revenue-spend hypothesis (Friedman, 1978; Buchanan and Wagner, 1977) contradicts this, asserting that tax revenues are a predecessor to government spending with the latter increasing as taxes go up or down: higher taxes will fuel government expenditures while tax cuts create an illusion that actually causes spending to jump. Both

hypotheses recognize a link between revenues and expenditures but differ in the proposed causality domains: the spend-tax sees spending as the cause while the revenue-spend sees taxation as the source.

A contrasting theory on public finance, Lerner's functional finance (1943), holds that the regulatory interest of economic stability shall receive precedence over expenditure and taxation to an extent. The tenets of classical theory would be against contracting debts. Smith (1937) saw public debt as a terrible detriment, while Mill (1979) found it a good, as the interest being paid on sufficiently settled foreign wealth, is able to divert investment expenditures into the real domestic economy.

In contrast to classical views on public debt, Keynesian theory (1936) postulates that increasing government expenditure and cutting taxes will spur economic growth, especially in times of recession. Thus, contrary to Keynesianism, the revenue-spend hypothesis (Friedman 1978) argues that spending depends on revenue collection. In Mali, Keynesian principles are pertinent, as fiscal expansion usually gives way to debt accumulation, which lends itself to the displacement effect. However, in the face of fiscal constraints, the revenue-spend hypothesis considers revenue generation as a predecessor to expenditures, attaching competitiveness. Their presence seemingly illustrates the contrast among theorists rowing in one river of thought concerning dynamics of fiscal policies in developing countries.

## **Empirical Literature**

The relationship between government revenue, expenditure, and public debt has been widely examined with mixed findings. Iiyambo and Kaulihowa (2020) found borrowing drives spending in Namibia, supporting the tax-spend hypothesis—similar to findings in Kenya (Kiminyei, 2014), Nigeria (Uguru, 2016), and Jordan (Alawneh, 2017). Others highlight factors like corruption (Del Monte & Pennacchio, 2020), fiscal transparency (Roth et al., 2022), and weak tax systems (Mose et al., 2024). Debt-growth dynamics vary across contexts (Butkus et al., 2021; Berkeley et al., 2022; Alhamdany et al., 2025). While some support investment-led spending (Awoyemi, 2020; Bahaa, 2021), others emphasize institutional reforms (Menguy, 2024; Spyrakis & Kotsios, 2021). Given limited evidence for Mali, this study addresses the gap by analyzing its fiscal dynamics.

## Methodology

### ❖ Research Design

This methodology employs a quantitative approach utilizing the Vector Error Correction Method (VECM) to analyze the phenomena of government revenue, expenditure, and public debt in Mali for the period of 2000 to 2024. This period guarantees the reliability of data from institutions such as the IMF and World Bank and corresponds with some of the key fiscal reforms and economic developments in Mali. Further, it evaluates the impacts of recent shocks and resilience of Mali's economy, thereby giving optimum insights into fiscal policy and financial stability.

### ❖ Model Specification and Data Sources

The study employs a multiple linear regression approach, adapting the methodology of Abdulrasheed (2017), with public debt as an additional explanatory variable. This inclusion aligns with Favero and Giavazzi (2007), who emphasized the importance of public debt in policy analyses.

The inclusion of public debt is justified as government revenue alone is often inadequate to finance government expenditure. Public debt affects fiscal policy through interest payments, debt servicing, and principal repayments, making it a critical factor in understanding the dynamics of government fiscal sustainability.

Thus, the functional relationship among the variables is specified as:

$$GVTRV_t = f(GVTEX_t, PD_t)$$

Transforming variables into natural logarithms enhances model estimation by addressing heteroskedasticity, ensuring stationarity, and enabling elasticity interpretation. It stabilizes variance, maintains consistent statistical properties over time, and allows coefficients to reflect percentage changes, improving robustness and interpretability.

The log-linear regression model is specified as follows:

$$\ln(GVTRV_t) = \alpha + \beta_1 \ln(GVTEX_t) + \beta_2 \ln(PD_t) + \varepsilon_t$$

The model examines the relationship between government revenue ( $GVTRV_t$ ), expenditure ( $GVTEX_t$ ), and public debt ( $PD_t$ ) over 2000–2024 using World Bank and IMF data.  $GVTRV_t$  is the dependent variable, reflecting government income, while  $GVTEX_t$  and  $PD_t$  represent spending and borrowing. The intercept ( $\alpha$ ) indicates baseline revenue, and coefficients  $\beta_1$  and  $\beta_2$  measure revenue elasticity concerning expenditure and debt. The error term ( $\epsilon_t$ ) captures unobserved factors. This framework analyzes fiscal dynamics without relying on original data sources, ensuring consistency across periods.

### **Estimation Techniques**

The study applied the Classical Linear Regression Model (CLRM) with Ordinary Least Square estimation, augmented with descriptive statistics and time-series tests such as unit root and cointegration tests. Then, the appropriate model was determined. Short-run and long-run relationships among public debt, government revenue, and expenditure were captured with a Vector Error Correction Model (VECM), which was further subjected to Granger causality tests for the directional relationships. Data analysis was conducted using Stata.

### **Stationary and Unit Root Tests**

A time series is stationary if its mean and variance remain constant (Gujarati, 2004). To prevent spurious results, stationarity was tested using the Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) tests. Both tests check for unit roots, with the null hypothesis assuming non-stationarity. Variables were tested at levels and first differences, and if integrated of order one [ $I(1)$ ], cointegration analysis followed.

### **Cointegration Test**

Once the unit root tests results confirmed the non-stationarity of variables at levels, cointegration analysis was considered, which implies the existence of long-run equilibrium relationship among the variables that have been incorporated in a model (Gujarati, 2004).

### **Johansen cointegration**

The trace and maximum eigen value Johansen tests were used. The null hypothesis is rejected if the probability value (p-value) under both trace and maximum

eigenvalue tests are less than 5% (0.05) significance level and the statistics for both tests are greater than the critical value at 5% level of significance, concluding that the variables are cointegrated. The optimal lag length was also determined based on the model with the lowest Akaike Information Criterion (AIC) or Schwarz Information Criterion (SC) values.

### Vector Error Correction Model (Vecm)

The study establishes a Vector Error Correction Model-VECM as applied to a log transformed level of government revenue, expenditure, and public debt as required by the cointegration framework of Engle and Granger (1987). Indeed, the Johansen Cointegration Test (1988, 1991) reveals that these variables exhibit a long-run equilibrium and thus, allows a VECM instead of first-differencing in order not to lose the long-term relationships induced by the underlying model. This vector autoregression with cointegration constraints is expected to capture short-term dynamics as well as long-run adjustments and, of course, the speed of reversion to equilibrium so determined by the error correction term.

The VECM equation takes the form:

$$\Delta Y_t = \alpha + \sum_{i=1}^p \Gamma_i \Delta Y_{t-i} + \Pi Y_{t-i} + \varepsilon_t$$

$Y_t$  is the Vector of log-transformed level values of the dependent variables (government revenue, government expenditure, public debt)?  $\Delta Y_t$  First differences of the variables, capturing short-run adjustments.  $\alpha$  represents Vector of constant terms (intercepts). While this  $\sum_{i=1}^p \Gamma_i \Delta Y_{t-i}$  captures the short-run dynamics, showing how past changes in revenue, expenditure, and debt influence current values.  $\Pi Y_{t-i}$  includes the error correction term (ECT), which measures how the system adjusts back to long-run equilibrium.  $\varepsilon_t$  Vector of white noise error terms, representing unexplained variations. The Error Correction Term (ECT) is analyzed to determine the speed at which fiscal disequilibria adjust toward long-run equilibrium.

### Granger Causality Tests

In this study, the Chow test (Chi-square) in the Vector Error Correction Model has been used to determine causality among revenues, expenditures, and debts. The Granger causality test establishes relationships in time series data, whereas VECM modifies Granger causality by allowing the possibility of the variables being

cointegrated and answer short- and long-term dynamics. Wald's test evaluates the significance of coefficients in testing for Granger causality, similar to the modified test of Toda and Yamamoto (1995), which resolves issues with integration orders.

To test whether government revenue Granger-causes government expenditure, we estimate:

$$GVTEX_t = \alpha_0 + \sum_{i=1}^p \beta_i GVTEX_{t-i} + \sum_{i=1}^p \gamma_i GVTRV_{t-i} + \sum_{i=1}^p \delta_i PD_{t-i} + \varepsilon_t$$

This equation tests whether past values of government revenue (GVTRV) influence government expenditure (GVTEX). Here,  $GVTEX_t$  represents government expenditure at time  $t$ . while  $\alpha_0$  is the constant term. The sum  $\sum_{i=1}^p \beta_i GVTEX_{t-i}$  reflects how previous government spending influences current expenditure, suggesting that spending patterns may follow historical trends. Similarly,  $\sum_{i=1}^p \gamma_i GVTEX_{t-i}$  captures the impact of past revenue on present expenditure. If the coefficients  $\gamma_i$  are statistically significant, it indicates that past revenue plays a role in shaping government spending, aligning with the tax-spend hypothesis—the idea that governments first collect revenue before determining expenditure levels. Additionally,  $\sum_{i=1}^p \delta_i PD_{t-i}$  accounts for the influence of past public debt on current spending.  $\varepsilon_t$  is the error term, accounting for factors not included in the model.

If the null hypothesis  $H_0 : \gamma_1 = \gamma_2 = \dots = \gamma_p = 0$  is rejected, it implies that government revenue Granger-causes government expenditure, meaning tax revenues strongly influence government spending decisions.

To test whether government expenditure Granger-causes government revenue, we estimate:

$$GVTRV_t = \alpha_1 + \sum_{i=1}^p \lambda_i GVTRV_{t-i} + \sum_{i=1}^p \theta_i GVTEX_{t-i} + \sum_{i=1}^p \omega_i PD_{t-i} + \varepsilon_t$$

This equation investigates the reverse relationship: Does government expenditure influence government revenue? Here,  $GVTRV_t$  represents government revenue at time  $t$ . While  $\alpha_1$  is the constant term. The sum  $\sum_{i=1}^p \lambda_i GVTRV_{t-i}$  captures the role of past revenue levels on current revenue, ensuring that past fiscal policies and trends are accounted for. Similarly,  $\sum_{i=1}^p \theta_i GVTEX_{t-i}$  measures the effect of past government spending on current revenue. If the coefficients  $\theta_i$  are statistically significant, it indicates that government spending affects future revenue, supporting

the spend-tax hypothesis—where governments spend first and then adjust taxes accordingly to finance expenditure. Additionally,  $\sum_{i=1}^p \omega_i PD_{t-i}$  reflects the role of public debt in shaping revenue collection efforts.  $\varepsilon_t$  represents the error term.

If the null hypothesis  $H_0: \theta_1 = \theta_2 = \dots = \theta_p = 0$  is rejected, government expenditure Granger-causes government revenue, suggesting that tax policies respond to changes in government spending rather than the other way around.

To test whether public debt Granger-causes government revenue, we estimate:

$$GVTRV_t = \alpha_2 + \sum_{i=1}^p \lambda_i GVTRV_{t-i} + \sum_{i=1}^p \tau_i PD_{t-i} + \sum_{i=1}^p \psi_i GVTEX_{t-i} + \varepsilon_t$$

This equation examines whether past levels of public debt impact government revenue collection. Here,  $GVTRV_t$  represents government revenue at time  $t$ . While  $\alpha_2$  is the constant term. The sum  $\sum_{i=1}^p \lambda_i GVTRV_{t-i}$  accounts for the role of past revenue on current revenue. Similarly,  $\sum_{i=1}^p \tau_i PD_{t-i}$  captures the effect of past debt levels on revenue collection efforts. If the coefficients  $\tau_i$  are statistically significant, it implies that increasing public debt influences government revenue generation, possibly due to higher taxation imposed to service debt. Additionally,  $\sum_{i=1}^p \psi_i GVTEX_{t-i}$  controls for the impact of past government expenditure on current revenue.  $\varepsilon_t$  represents the error term.

If the null hypothesis  $H_0: \tau_1 = \tau_2 = \dots = \tau_p = 0$  is rejected, it suggests that public debt Granger-causes government revenue, meaning that growing debt influences tax revenue collection, possibly due to fiscal adjustments such as higher tax rates or improved collection efficiency.

To test whether public debt Granger-causes government expenditure, we estimate:

$$GVTEX_t = \alpha_3 + \sum_{i=1}^p \beta_i GVTEX_{t-i} + \sum_{i=1}^p \sigma_i PD_{t-i} + \sum_{i=1}^p \rho_i GVTRV_{t-i} + \varepsilon_t$$

This equation examines whether past debt accumulation influences government spending decisions. Here,  $GVTEX_t$  represents government expenditure at time  $t$ . While  $\alpha_3$  is the constant term. The sum  $\sum_{i=1}^p \beta_i GVTEX_{t-i}$  captures the role of past expenditure on current expenditure, ensuring fiscal trends are considered. Similarly,  $\sum_{i=1}^p \sigma_i PD_{t-i}$  measures the impact of past debt levels on government spending. If the coefficients  $\sigma_i$  are statistically significant, it suggests that higher debt levels drive increased government spending, possibly due to interest payments or expansionary fiscal policies fueled by borrowing. Additionally,  $\sum_{i=1}^p \rho_i GVTRV_{t-i}$

controls for the influence of past government revenue on expenditure.  $\varepsilon_t$  represents the error term.

If the null hypothesis  $H_0 : \sigma_1 = \sigma_2 = \dots = \sigma_p = 0$  is rejected, it implies that public debt Granger-causes government expenditure, meaning that rising debt levels directly influence how much the government spends.

### Model Diagnostics

To ensure model robustness, diagnostic tests are conducted: the Autocorrelation Test (Lagrange Multiplier) checks for serial correlation, the Heteroskedasticity Test (White’s Test) assesses error variance stability, and the Normality Test (Jarque-Bera) verifies residual normality. These tests confirm whether assumptions hold or differ across models.

## Empirical Results and Discussions

### Descriptive Statistics

The Table 1 provides descriptive statistics for government revenue (GVTRV), expenditure (GVTEX) and public debt (PD) in emphasizing the main trends in fiscal management for Mali. Expenditure, on average (3.0679) exceeds revenue (2.7707), suggesting fiscal imbalances and deficit accumulation (3.5955). Public debt is the most volatile (SD=0.4308), proving that it fluctuates in any analysis. This usually occurs where policies are weak, or there is an external shock. Revenue, debt, and expenditures are generally positively skewed, although expenditures are slightly negatively skewed. Based on the findings of the Jarque-Bera test, they confirm an approximate normality, hence are fit for econometric modeling. Therefore, these findings speak well with effective debt and fiscal policies that would close up structural deficits.

**Table 1** Descriptive Statistics

Statistic	(GOVTRV)	(GOVTEX)	(PD)
Mean	2.7707	3.0679	3.5955
Median (50%)	2.7279	3.0634	3.642
Maximum	3.0865	3.2658	4.5054
Minimum	2.5014	2.740	2.8959

Standard Deviation	0.1677	0.1263	0.4308
Skewness	0.4969	-0.2268	0.1281
Kurtosis	2.2176	3.2465	2.3241
Jarque-Bera	1.667	0.2777	0.5442
Probability	0.4346	0.8704	0.7618
Sum	69.2675	76.6975	89.8875
Observations (n)	25	25	25

Source: Author's compilation

### Stationarity Test

The unit root tests (ADF and PP) were conducted at levels and first differences. A variable is I (1) if it becomes stationary after first differencing, I (2) if differenced twice, and I (0) if stationary at levels. The null hypothesis of a unit root is rejected when the p-value is below 0.05 or when t-statistics exceed critical values. Results confirm that all variables ( $\Delta \text{Ln}(\text{GVTRV})$ ,  $\text{NM}(\text{GVTEX})$ , and  $\text{NM}(\text{DEBT})$ ) are I (1), ensuring compatibility for cointegration analysis and VECM estimation.

**Table 2 Unit root test results at levels-ADF and PP**

Variable	$\Delta \text{Ln}(\text{GVTRV})$	$\Delta \text{Ln}(\text{GVTEX})$	$\Delta \text{Ln}(\text{DEBT})$
Deterministic Terms	Intercept & Trend	Intercept & Trend	Intercept & Trend
ADF Test Statistic	-7.69	-7.218	-5.383
5% Critical Value	-3.600	-3.600	-3.600
P-Value	0.0000	0.0000	0.0000
PP Test Statistic	-7.731	-7.463	-5.457
5% Critical Value	-3.600	-3.600	-3.600
P-Value	0.0000	0.0000	0.0000
Order of Integration	I(1)	I(1)	I(1)

Source: Author's compilation and data analysis using Stata.

## Cointegration Test

### ❖ The Johansen Cointegration Test

The Johansen cointegration test was conducted after determining the optimal lag length (AIC = 1). The trace and maximum eigenvalue tests confirm cointegration, rejecting the null hypothesis of no cointegration. This indicates a long-run relationship among public debt, government expenditure, and government revenue over the study period.

**Table 3 Lag Length Selection Criteria**

Sample: 2004 thru 2024 | Number of Observations = 21

Criterion	Lag 0	Lag 1	Lag 2	Lag 3	Lag 4
Log-Likelihood (LL)	39.023	56.6702	61.8427	65.6075	78.7578
Likelihood Ratio (LR)	-	35.294	10.345	7.5296	26.301*
Degrees of Freedom (df)	-	9	9	9	9
P-Value	-	0.000	0.323	0.582	0.002
Final Prediction Error (FPE)	6.5e-06	2.9e-06*	4.4e-06	8.7e-06	8.5e-06
AIC	-3.43077	-4.2543*	-3.88978	-3.39119	-3.78645
HQIC	-3.39838	-4.12477*	-3.66309	-3.06735	-3.36546
SBIC	-3.28155	-3.65743*	-2.84526	-1.89902	-1.84663

\* optimal lag

Endogenous: *log\_TotalGovernmentRevenue log\_TotalExpenditure*

Exogenous: *\_cons*

Source: Author.

**Table 4 Johansen Cointegration Test Results**

Null Hypothesis (H <sub>0</sub> )	Trace Statistic	5% Critical Value	Decision
No cointegration (r = 0)	34.2756	29.68	Reject H <sub>0</sub> (Cointegration Exists)

At most 1 cointegration ( $r \leq 1$ )	11.5894	15.41	Fail to Reject $H_0$
At most 2 cointegration ( $r \leq 2$ )	0.3816	3.76	Fail to Reject $H_0$
Null Hypothesis ( $H_0$ )	Max Eigenvalue Statistic	5% Critical Value	Decision
No cointegration ( $r = 0$ )	22.6862	20.97	Reject $H_0$ (Cointegration Exists)
At most 2 cointegration ( $r \leq 1$ )	11.2078	14.07	Fail to Reject $H_0$
At most 2 cointegration ( $r \leq 2$ )	0.3816	3.76	Fail to Reject $H_0$

Source: Author.

### ❖ Vector Error Correction

**Table 5** Vector Error Correction Model (VECM)

#### 1. Error Correction Term (ECT)

Dependent Variable	ECT Coefficient	Std. Error	z-Statistic	P-Value	Adjustment Speed
$\Delta \text{Ln}(\text{GVTRV})$ (Government Revenue)	-0.1505	0.2324	-0.65	0.517	Not significant
$\Delta \text{Ln}(\text{GVTEX})$ (Government Expenditure)	0.6998	0.2028	3.45	0.001	Significant, adjusts to equilibrium
$\Delta \text{Ln}(\text{DEBT})$ (Public Debt)	0.5669	0.5701	0.99	0.320	Not significant

Source: Author.

**2. Long-Run Cointegrating Relationship**

Variable	Coefficient	Std. Error	z-Statistic	P-Value	Significance
Ln(GVTEX) (Government Expenditure)	-1.3017	0.1291	-10.08	0.000	Significant
Ln(DEBT) (Public Debt)	-0.0100	0.0368	-0.27	0.785	Not significant
Constant	1.2539	-	-	-	-

Source: Author.

The Johansen normalization restriction estimates the long-run equilibrium relationship as:

$$\text{Ln(GVTR)} = 1.3017 \text{ Ln(GVTEX)} + 0.0100 \text{ Ln(DEBT)} - 1.2539$$

**3. Short-Run Dynamics (Chi<sup>2</sup> Tests)**

Equation	R <sup>2</sup> (Goodness of Fit)	Chi <sup>2</sup> (P-Value)	Short-Run Significance
$\Delta\text{Ln(GVTRV)}$ (Revenue)	0.0732	0.4195	No significant short-run effects
$\Delta\text{Ln(GVTEX)}$ (Expenditure)	0.3598	0.0021	Significant short-run effects
$\Delta\text{Ln(DEBT)}$ (Public Debt)	0.0496	0.5632	No significant short-run effects

Source: Author.

The R<sup>2</sup> values and Chi-square statistics measure how well short-run changes in revenue, expenditure, and debt explain each other.

The Vector Error Correcting Model (VECM), lagged by 1, is aptly specified, with an AIC value of -3.9767. The absence of autocorrelation confirmed by the LM test ( $p >$

0.05) is further evidence of the model's statistical validity. Analysis indicates that government revenue, expenditure, and public debt maintain a long-run cointegrating relationship with one another, consistent with Arestis et al.'s (2004) and Baharumshah et al.'s (2017) findings. Government spending significantly reduces revenue (-1.3017,  $p = 0.000$ ), in support of Fatás and Mihov (2003) on the fiscal burden of such excessive expenditures. However, public debt holds no significant revenue contribution (-0.0100,  $p = 0.785$ ), affirming Reinhart and Rogoff's (2010) findings that debt does not usually result in growth for the fiscal. Expenditure adjusts to fiscal imbalances ( $ECT = 0.6998$ ,  $p = 0.001$ ), while revenue and debt remain unresponsive, mirroring Afonso and Jalles (2013) spending-driven adaptations. Short-run instability is also caused by these expenditures, per Claeys et al. (2012). It is posited that inefficient revenue mobilization or deficit financing is likely to be the source of the negative expenditure-revenue link (Gupta et al. 2005). Public debt has no impact, indicating poor allocation (Cecchetti et al. 2011). As indicated by Gemmill et al. (2011), structural factors shape fiscal outcomes.

#### ❖ Granger Causality

A Granger causality test using the Wald test within a VECM framework examined short-run causal links among government revenue, expenditure, and debt. Three hypotheses were tested: (1) whether expenditure and debt influence revenue, (2) whether revenue and debt drive spending, and (3) whether revenue and spending affect debt levels. The test assessed the predictive significance of past values, revealing fiscal interdependencies.

**Table 6 Granger Causality Test Results (Wald Test in VECM)**

Null Hypothesis ( $H_0$ )	Chi <sup>2</sup> Statistic	P-Value	Decision (5% level)	Interpretation
Government revenue is not Granger-caused by expenditure & debt	34.62	0.0000	Reject $H_0$	Expenditure & debt Granger-cause revenue

Government expenditure is not Granger-caused by revenue & debt	34.62	0.0000	Reject $H_0$	Revenue & debt Granger-cause expenditure
Public debt is not Granger-caused by revenue & expenditure	34.62	0.0000	Reject $H_0$	Revenue & expenditure Granger-cause debt

Source: Author

There is a two-way relationship between government revenue and expenditure, which may imply both tax-spend and spend-tax hypotheses, according to Granger causality results. This agrees with Yashobanta (2012), who emphasized fiscal interdependence in India; these outcomes also show that Granger causes public debt by both revenues and expenditures. Hence, fiscal deficits in Mali are mainly financed through borrowings. This finding agrees with Eberhardt and Presbitero (2015) and Chudik et al. (2017), which warn about the dangers of borrowing in low-income economies. Revenue and expenditure dependence show that there is reactive fiscal policy in Mali-the same is also observed by Ugwuanyi et al. (2017) in sub-Saharan Africa. Given the reliance on borrowing, effective debt management is crucial to avoid sustainability risks, as underscored by Gunduz (2017).

❖ **Diagnostic Test**

**Table 7 Diagnostic tests for autocorrelation, heteroskedasticity, and Normality.**

Diagnostic Test	Test Statistic	Degrees of Freedom (DF)	P-Value	Decision (5% level)	Conclusion
Autocorrelation (LM Test - Lag 1)	11.0798	9	0.2703	Fail to reject $H_0$	No autocorrelation

Autocorrelation (LM Test - Lag 2)	10.8783	9	0.2841	Fail to reject Ho	No autocorrelation
Heteroskedasticity (White Test)	8.04	9	0.5301	Fail to reject Ho	No heteroskedasticity
Normality (Total Test)	15.81	13	0.2595	Fail to reject Ho	Residuals are normally distributed
Normality (Skewness Test)	4.42	3	0.2200	Fail to reject Ho	No strong skewness
Normality (Kurtosis Test)	3.36	1	0.0670	Slightly rejects Ho	Possible excess kurtosis, but not severe

Source: Author

## Conclusions and Recommendations

This study analyzes Mali's fiscal dynamics (2000–2024) using a Vector Error Correction Model (VECM) and Granger causality tests. Findings indicate a stable long-run relationship between revenue, expenditure, and debt, with inefficient tax collection and deficit financing undermining fiscal health. Government spending negatively impacts revenue (-1.3017 elasticity), and debt sustainability remains uncertain.

Short-term adjustments show expenditure responds quickly to imbalances, while revenue and debt remain sluggish. Policy recommendations focus on formalizing the informal sector, digitizing tax collection, optimizing key sectors (gold mining, telecom), and harmonizing WAEMU tax policies. Expenditure reforms emphasize productive investments, outcome-based budgeting, and reducing non-essential spending. Strengthened debt management, concessional borrowing, and fiscal rules are crucial for stability.

Additionally, integrating climate resilience into fiscal planning through environmental taxes and green financing will bolster sustainability. Comprehensive reforms in tax efficiency, spending rationalization, and debt control are essential for ensuring Mali's long-term fiscal stability and economic resilience.

### **Data Availability**

The datasets used in this study are publicly available from the International Monetary Fund (IMF) and the World Bank databases. The data supporting the findings of this research can be accessed through the respective websites of these institutions. Additional details or processed data used in the analysis are available from the corresponding author upon reasonable request.

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